

Survival analysis

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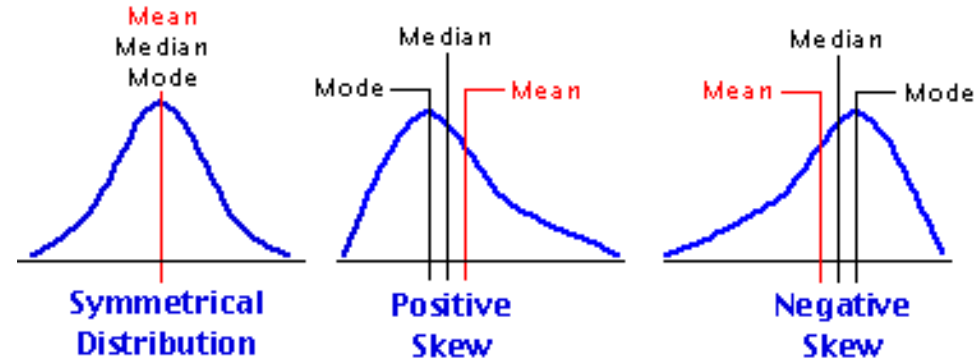
Survival analysis

- Time to **event data**.
- **Censoring**.
- **Survivor function – Kaplan-Meier plot**.
- **Log-rank test**.
- **Hazard function and Hazard ratio** .

Time to event data: examples

- Time to death.
- Time to progression of cancer.
- Time to development of diabetes.
- Time to recovery from diarrhea.
- Time to event data typically collected in
 - cohort studies (time between study baseline and event of interest).
 - clinical trials (time between randomisation and event of interest).
- Also known as **survival data**.

Features of time to event data

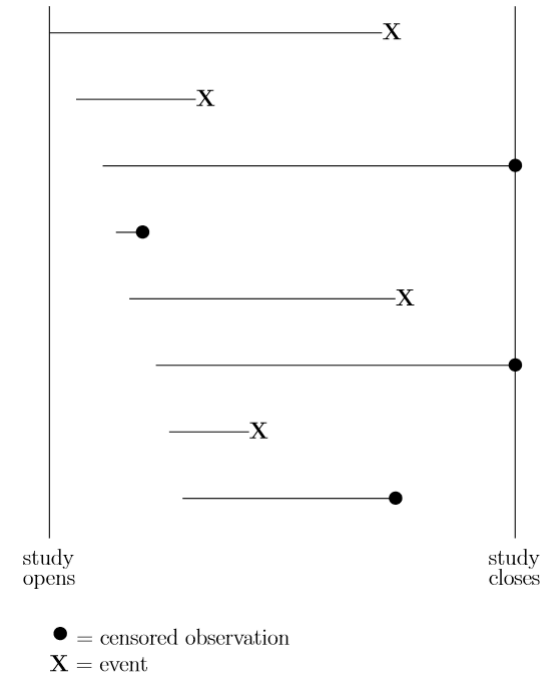


- Non-negative values.
- Not normally distributed (usually positively skewed).
- Event not usually observed for all individuals during the study.
- An observation is **censored** if individual does not experience event during the study.
- **Censoring time:** time from baseline/randomisation until latest date at which individual is known to be still alive and event-free.

Censoring

- Definition: Event of interest not observed for all individuals.
- **Fixed censoring:** event has not occurred when study has ended or data analysis is performed.
- **Loss to follow-up:** individual has been lost to follow-up (e.g. he/she no longer wishes to take part in study).

Illustration of survival data



- Survival analysis methods make use of information from censored observations.
- Assume censoring is **non-informative**, i.e. if an individual is censored, his/her subsequent risk of the event of interest is unaffected.

Example of time to event data

Weeks to death or censoring (*) in 20 adults with recurrent astrocytoma:

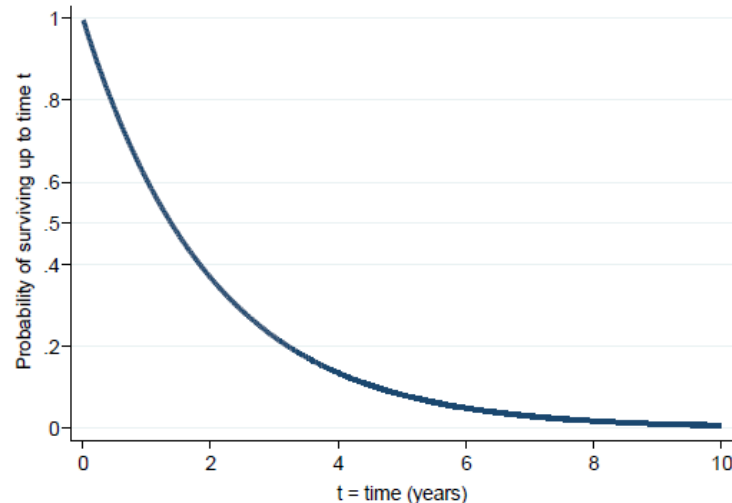
6	13	21	30	31*	37	38	47*	49	50
63	79	80*	82*	82*	86	98	149	202	219

ID	death	weeks
1	1	6
2	1	13
3	1	21
4	1	30
5	0	31
6	1	37
7	1	38
8	0	47
9	1	49
10	1	50
11	1	63
12	1	79
13	0	80
14	0	82
15	0	82
16	1	86
17	1	98
18	0	149
19	1	202
20	1	219

Aims of survival analysis

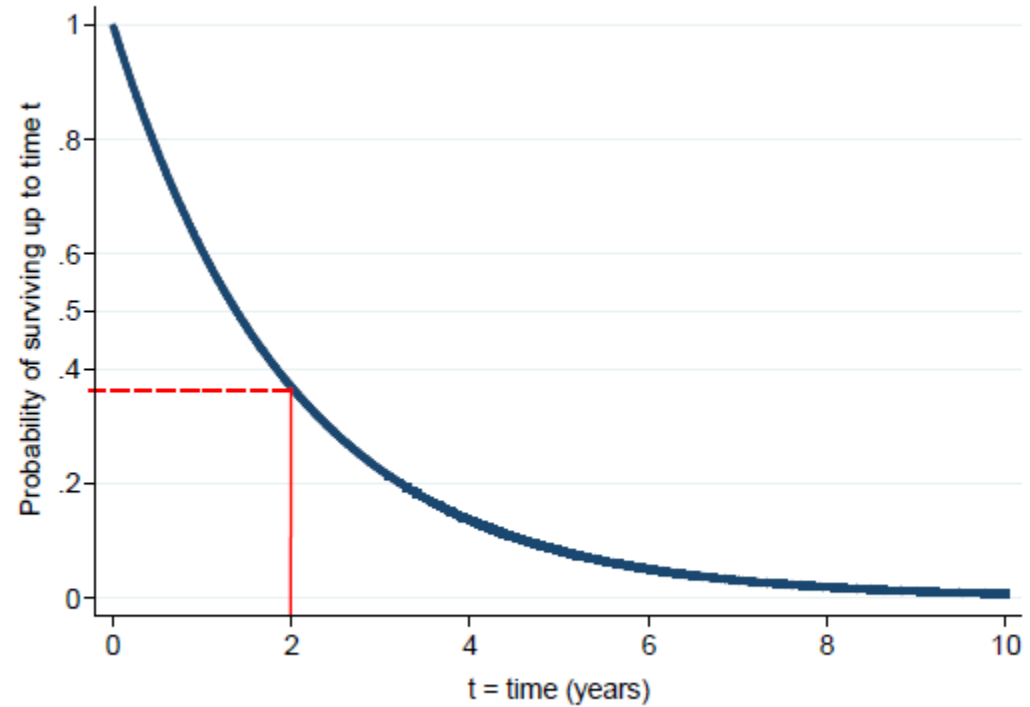
- To estimate probability of not experiencing event of interest (not dying = “surviving”) over any given time period (e.g. 5 year survival rate).
- To compare overall survival experience between different groups of individuals (e.g. between groups in a randomised clinical trial).
- **Survivor function:** Probability of not experiencing event of interest (“surviving”) up to time t .

Example:



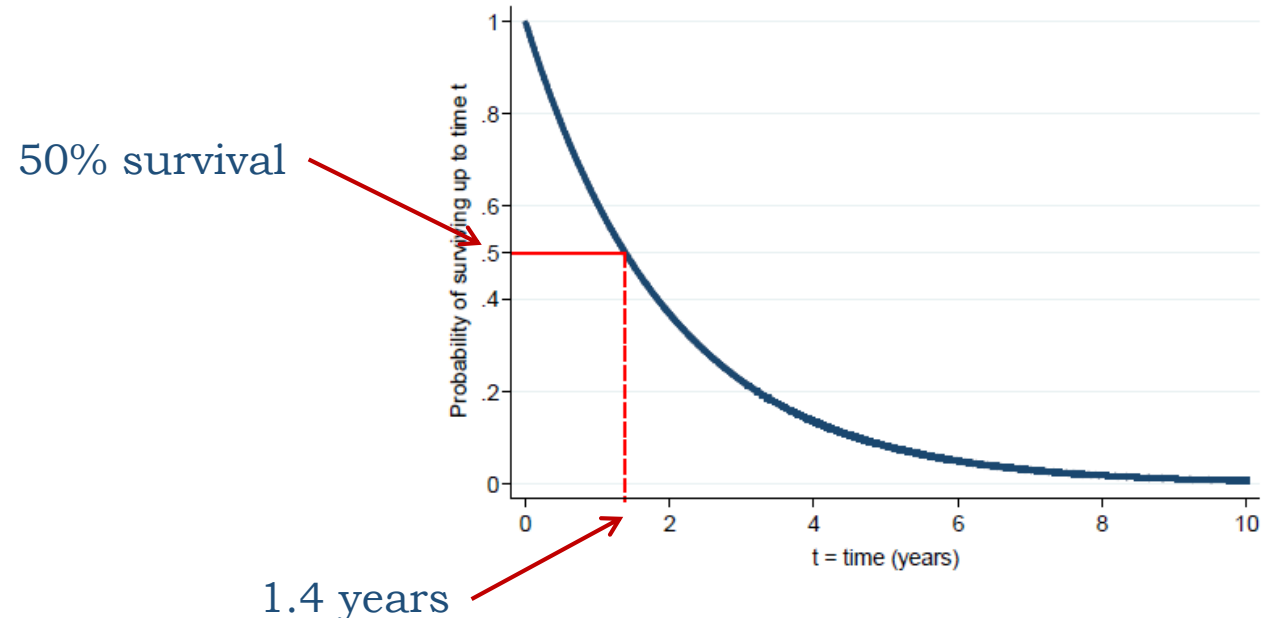
Estimating a survival rate

- Probability of surviving up to 2 years = 0.37.



Median survival time

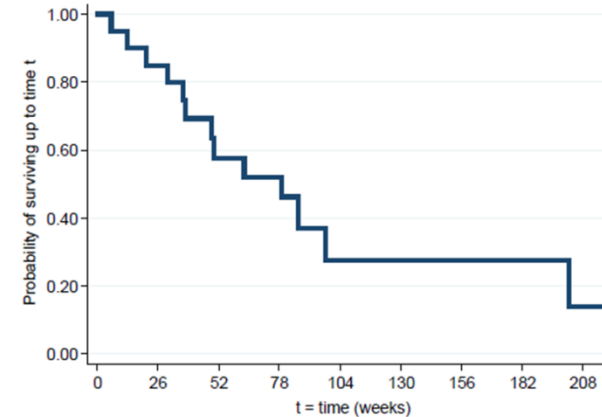
- It is the time (expressed in months or years) when half the patients are expected to be alive. It means that the chance of surviving beyond that time is 50%.
- Median survival time = 1.4 years, since the probability of surviving up to 1.4 years is 0.5.



Kaplan-Meier (KM) estimation of survivor function

First death

6	13	21	30	31*	37	38	47*	49	50
63	79	80*	82*	82*	86	98	149	202	219



- **20** individuals in study at $t=0$.
- First death at $t=6$ weeks.
- No individuals censored before $t=6$.
- Probability of death for each individual: $1/20=0.05$
- Therefore probability of surviving beyond $t=6$ is $(1-0.05)=0.95=19/20$.

Weeks in follow-up (t)	N at risk at time t	N of deaths at time t	Prob. of death at time t	Prob. of no death at time t	Prob. of surviving up to and including time t
0	20	0	0	1	1
6	20	1	0.05	0.95	$1 \times 0.95 = 0.95$

“Risk set” at time t

$1/20$

$19/20$

K-M estimation of survivor function

Second death

	13	21	30	31*	37	38	47*	49	50
63	79	80*	82*	82*	86	98	149	202	219

- **19** individuals in study between $t=6$ and $t=13$.
- Second death at $t=13$.
- No individuals censored between $t=6$ and $t=13$.
- Probability of death for each individual: $\frac{1}{19} = 0.053$ $\frac{19}{20}$ $\frac{18}{19}$
- Therefore probability of surviving beyond $t=13$ is $0.95 \times 0.947 = 0.90$.
 - with $0.95 = (1 - (1/20))$ and $0.947 = (1 - (1/19))$

Weeks in follow-up (t)	N at risk at time t	N of deaths at time t	Prob. of death at time t	Prob. of no death at time t	Prob. of surviving up to and including time t
6	20	1	0.05	0.95	0.95
13	19	1	0.053	0.947	$0.95 \times 0.947 = 0.90$

$$\frac{1}{19}$$

$$1 - (1/19) = 18/19$$

K-M estimation of survivor function

Third and fourth death

		21	30	31*	37	38	47*	49	50
63	79	80*	82*	82*	86	98	149	202	219

- **18** individuals in study between $t=13$ and $t=21$.
- Probability of death for each individual: **$1/18=0.056$**
- Probability of surviving beyond $t=21$ is **$0.90 \times (1-(1/18)) = 0.85$** . From $t=13$: 0.95×0.947
- **17** individuals in study between $t=21$ and $t=30$.
- Probability of death for each individual: **$1/17=0.059$**
- Probability of surviving beyond $t=30$ is **$0.85 \times (1-(1/17)) = 0.80$** .

Weeks in follow-up (t)	N at risk at time t	N of deaths at time t	Prob. of death at time t	Prob. of no death at time t	Prob. of surviving up to and including time t
13	19	1	$1/19= 0.053$	0.947	0.90
21	18	1	$1/18= 0.056$	0.944	0.85
30	17	1	$1/17= 0.059$	0.941	0.80

K-M estimation of survivor function

Fifth and sixth death

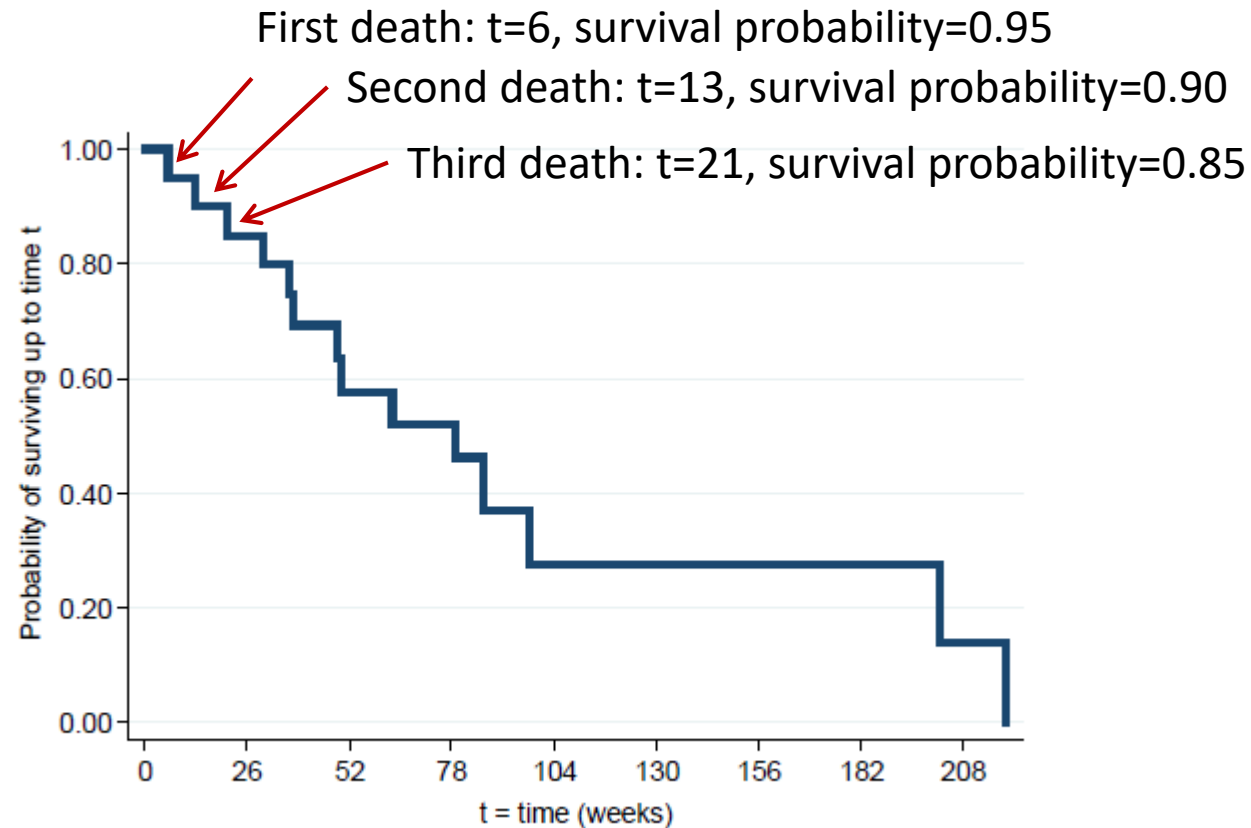
				31*	37	38	47*	49	50
63	79	80*	82*	82*	86	98	149	202	219

- **16** individuals in study between $t=30$ and $t=31$.
- 1 individual censored at $t=31$.
- **Probability of surviving beyond $t=31$ remains at 0.80.**
- **15** individuals in study between $t=31$ and $t=37$.
- Probability of surviving beyond $t=37$ is **$0.80 \times (1 - (1/15)) = 0.747$.**

Weeks in follow-up (t)	N at risk at time t	N of deaths at time t	Prob. of death at time t	Prob. of no death at time t	Prob. of surviving up to and including time t
30	17	1	0.059	0.941	0.80
31	16	0	0	1	$0.80 \times 1 = 0.80$
37	15	1	$1/15 = 0.067$	0.933	$0.80 \times 0.933 = 0.747$

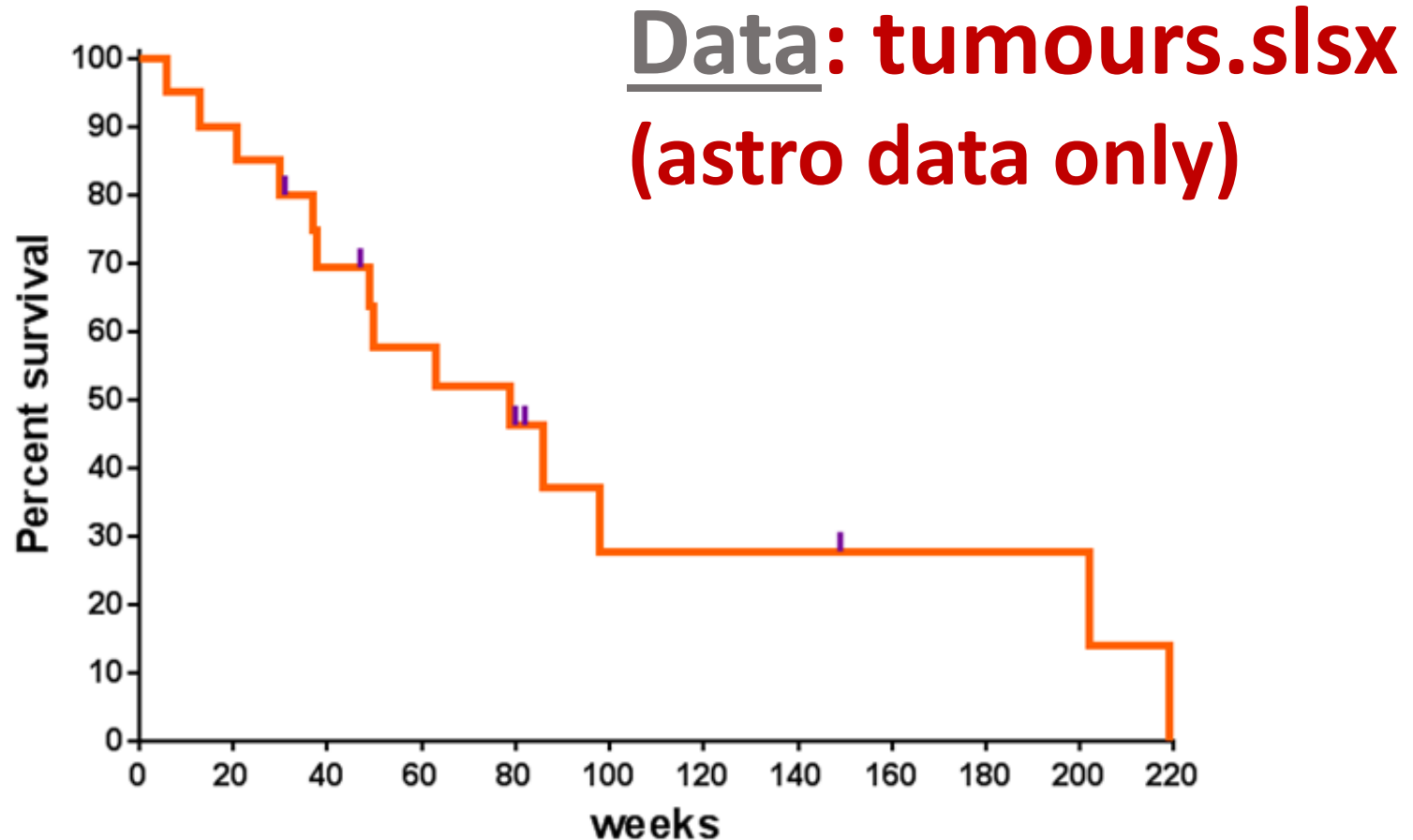
K-M plot of survivor function

- Continue these calculations until reaching the longest event time.
- K-M plot drawn as a step function:



K-M plot of survivor function

- Add ticks to indicate where censoring occurred.



Comparing 2 groups

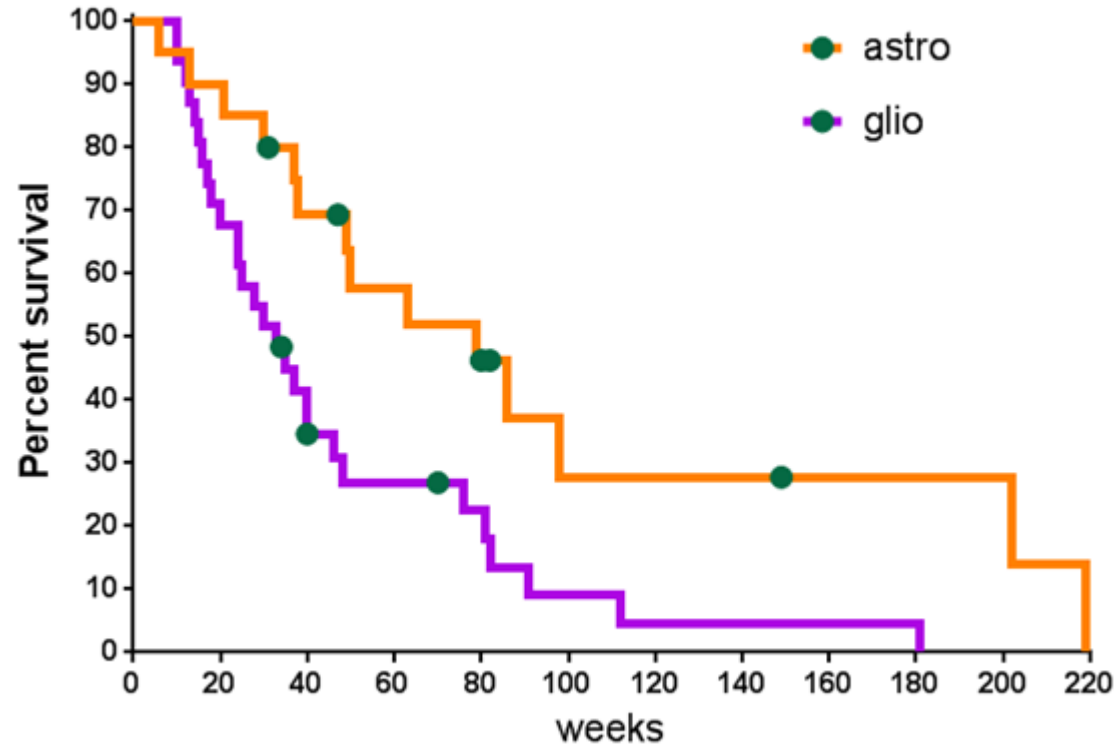
- Weeks to death or censoring (*) in **20 adults** with recurrent astrocytoma:

6	13	21	30	31*	37	38	47*	49	50
63	79	80*	82*	82*	86	98	149	202	219

- Weeks to death or censoring (*) in **31 adults** with recurrent glioblastoma:

10	10	12	13	14	15	16	17	18	20
24	24	25	28	30	33	34*	35	37	40
40	40*	46	48	70*	76	81	82	91	112
181									

K-M plot of survivor function by tumour type



- Survival chances appear better in individuals with astrocytoma than with glioblastoma, but is the **difference between groups statistically significant?**

Comparing 2 samples

- Could compare **median survival time**, or **probability of surviving** up to any particular time.
- Better to use a test which compares survivor functions over whole follow-up period.
- **Log rank test:** tests null hypothesis of no difference between samples in probability of an event (death in this example) at any time point during follow-up.
- **Log rank test statistic:**
 - based on calculating expected number of events that would occur under null hypothesis at each event time, and comparing to observed number of events.
 - under null hypothesis has a Chi^2 distribution with 1 degree of freedom.

Log rank test to compare 2 groups

Astro	Death (=1)	Glio	Death (=1)
6	1	10	1
13	1	10	1
21	1	12	1
30	1	13	1
31	0	14	1
37	1	15	1
38	1	16	1
47	0	17	1
49	1	18	1
50	1	20	1
63	1	24	1
79	1	24	1
80	0	25	1
82	0	28	1
82	0	30	1
86	1	33	1
98	1	34	0
149	0	35	1
202	1	37	1
219	1	40	1
		40	1
		40	0
		46	1
		48	1
		70	0
		76	1
		81	1
		82	1
		91	1
		112	1
		181	1

=14 deaths

=28 deaths

Week	Overall Observed Deaths	Expected Deaths – Astro	Expected Deaths – Glio	Observed Remainder – Astro	Observed Remainder – Glio
6	1/51	0.392157	0.607843	19	31
10	2/50	0.76	1.24	19	29
12					
13					
14					
15					
...					
Total (Expected)		Sum	Sum		
Total (Observed)		14	28		

20/51

31/51

(19/50)*2

(31/50)*2

Log rank test statistic has a Chi² distribution:

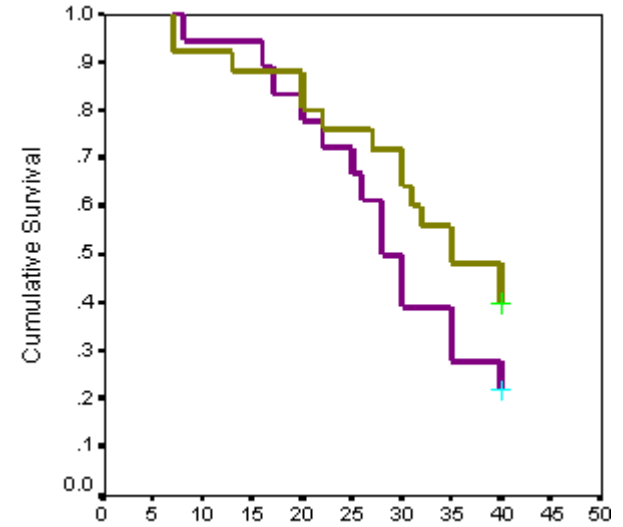
$$Z = \frac{\sum_{j=1}^J (O_{1j} - E_{1j})}{\sqrt{\sum_{j=1}^J V_j}}$$

Log rank test

- Unlikely to detect a difference between Groups if survivor functions cross over during follow-up.
- Assumes **non-informative censoring**
- Can be extended to compare more than 2 groups.

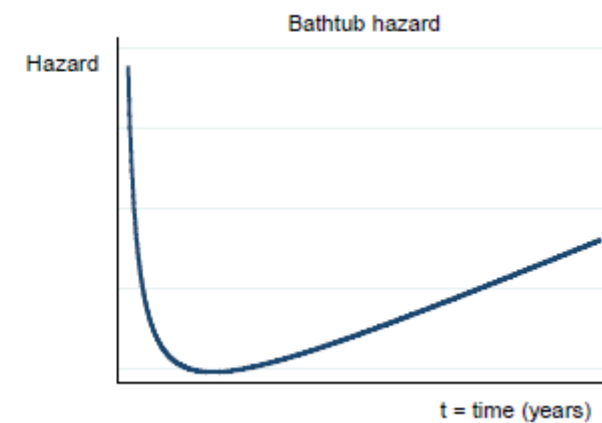
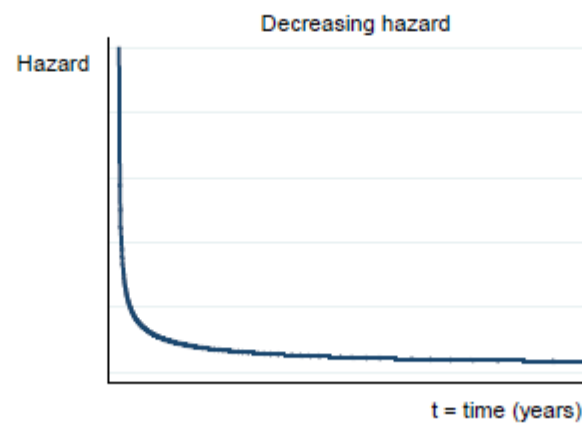
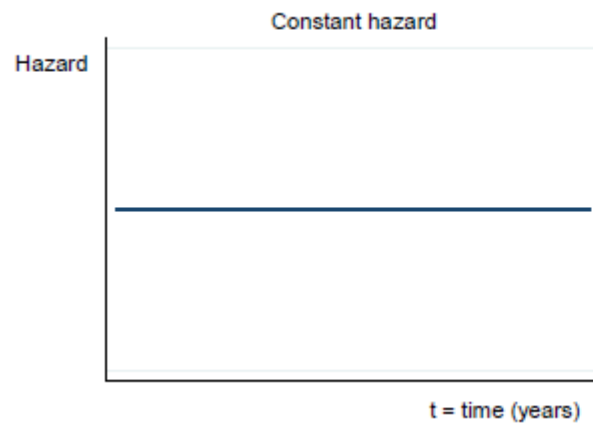
But

- Only provides a p-value, not an estimate of size of difference between groups or a confidence interval.
 - Estimate of size of difference = **Hazard Ratio**



Hazard function

- **Hazard** is defined as the slope of the survival curve :a measure of how rapidly subjects are dying.
- Hazard function describes how hazard varies over time.

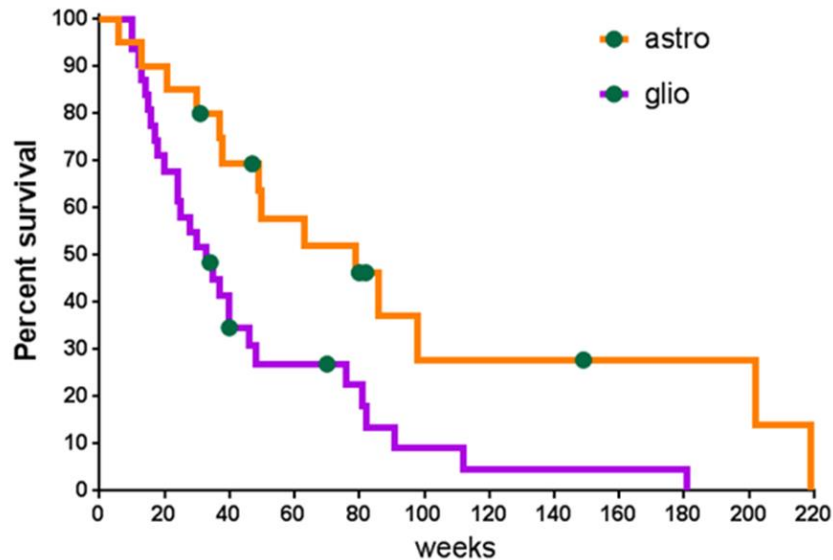


Hazard Ratio (HR) for comparing 2 samples

- Hazards may vary over time, but assume that **HR is constant over time.**
- The hazard ratio is not directly related to the ratio of median survival times.
- When comparing 2 groups (a and b):
 - observed events (deaths) in each group: **Oa** and **Ob**,
 - expected events (deaths) in each group: **Ea** and **Eb**,
 - assuming a null hypothesis of no difference in survival.
- **HR= (Oa/Ea)/(Ob/Eb)**
- No assumption is needed about shape of hazard functions or underlying distribution of time to event data.
- HR is obtained from **Cox regression**

Hazard Ratio (HR)

Data: tumours.xlsx



- **HR = 2.3** (95% CI [1.32;4.44])

- At any point in time, hazard (i.e. instantaneous rate) of dying in individuals with recurrent glioblastoma is **2.3 times** higher than in individuals with recurrent astrocytoma.

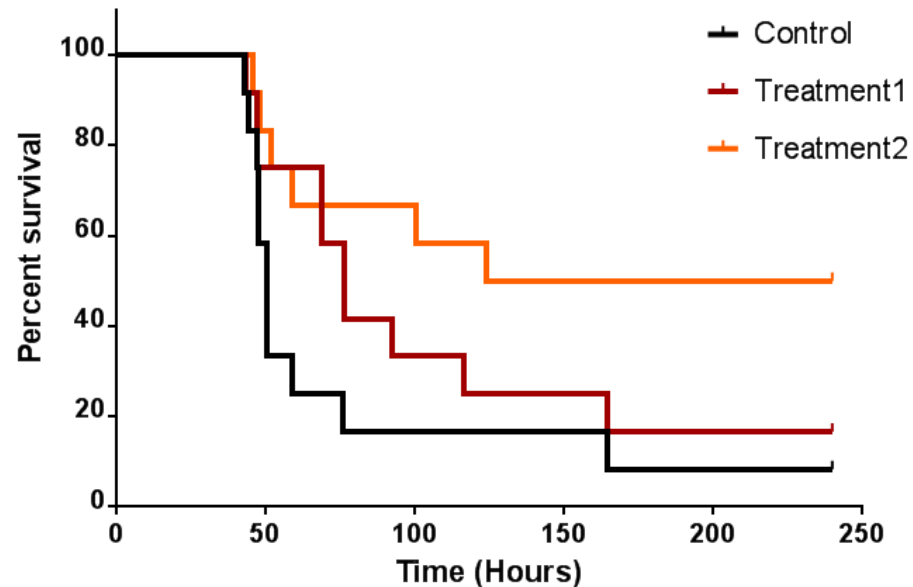
Survival Curve comparison		
1	Comparison of Survival Curves	
2		
3	Log-rank (Mantel-Cox) test	
4	Chi square	7.497
5	df	1
6	P value	0.0062
7	P value summary	**
8	Are the survival curves sig different? Yes	
9		
10	Gehan-Breslow-Wilcoxon test	
11	Chi square	5.828
12	df	1
13	P value	0.0158
14	P value summary	*
15	Are the survival curves sig different? Yes	
16		
17	Median survival	
18	astro	79
19	glio	33
20	Ratio (and its reciprocal)	2.394 0.4177
21	95% CI of ratio	1.26 to 4.547 0.2199 to 0.7934
22		
23	Hazard Ratio (Mantel-Haenszel)	
24	Ratio (and its reciprocal)	0.4132 2.42
25	95% CI of ratio	0.2194 to 0.7779 1.286 to 4.557
26		
27	Hazard Ratio (logrank)	
28	Ratio (and its reciprocal)	0.4341 2.304
29	95% CI of ratio	0.2367 to 0.7961 1.256 to 4.224
30		

Comparing more than 2 samples

- **Issue with GraphPad:** cannot compare more than 2 groups directly
 - As in: does not run post-hoc pairwise comparisons
- **So how do we do it?**
 - Step 1: All groups comparisons (equivalent omnibus step in ANOVA)
 - Step 2: Make all pairwise comparisons of interest
 - Step 3: Apply Bonferroni correction
- **Example dataset: Lung infection**
 - Mice are infected with *Streptococcus pneumoniae*
 - 3 groups: Control, treatment 1 and treatment 2

Comparing more than 2 groups

- Step 1: All groups comparisons



Comparison of Survival Curves	
Log-rank (Mantel-Cox) test (recommended)	
Chi square	7.112
df	2
P value	0.0286
P value summary	*
Are the survival curves sig different?	Yes
Logrank test for trend (recommended)	
Chi square	7.044
df	1
P value	0.0080
P value summary	**
Sig. trend?	Yes
Gehan-Breslow-Wilcoxon test	
Chi square	6.743
df	2
P value	0.0343
P value summary	*
Are the survival curves sig different?	Yes

- There is an overall difference in survival between the 3 groups but which group is different from which?

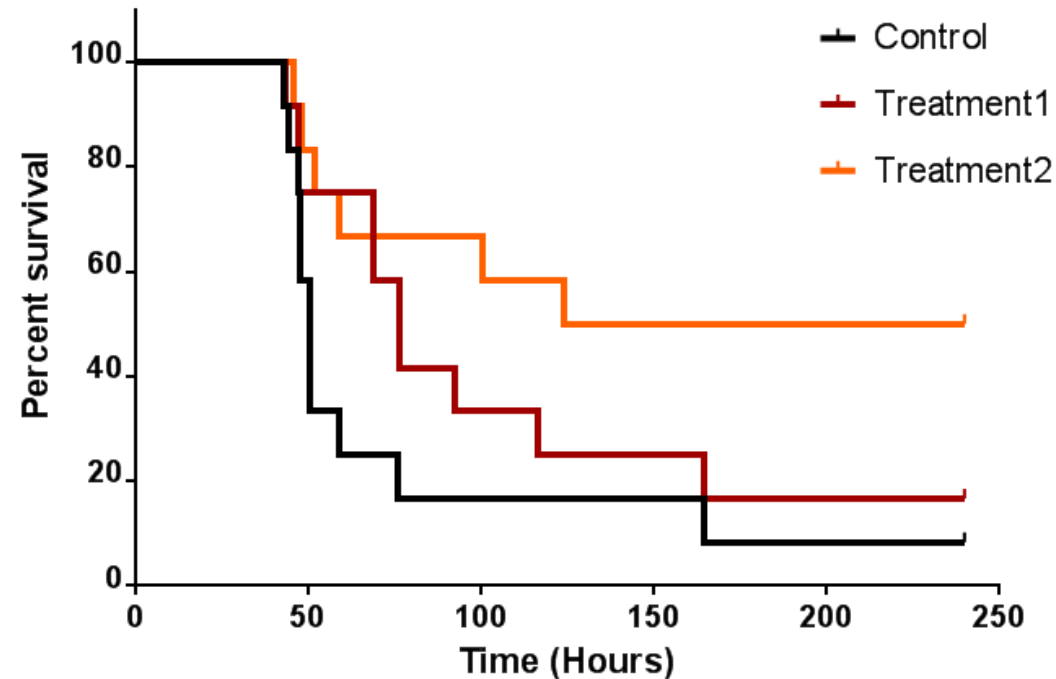
Comparing more than 2 groups

- Step 2: Make all pairwise comparisons of interest

Control vs. T1			Control vs. T2			T1 vs. T2		
Comparison of Survival Curves			Comparison of Survival Curves			Comparison of Survival Curves		
Log-rank (Mantel-Cox) test			Log-rank (Mantel-Cox) test			Log-rank (Mantel-Cox) test		
Chi square	1.800		Chi square	6.101		Chi square	2.214	
df	1		df	1		df	1	
P value	0.1798		P value	0.0135		P value	0.1367	
P value summary	ns		P value summary	*		P value summary	ns	
Are the survival curves sig different?	No	Adjusted p-value = 0.5394	Are the survival curves sig different?	Yes	Adjusted p-value = 0.0405	Are the survival curves sig different?	No	Adjusted p-value = 0.4101
Gehan-Breslow-Wilcoxon test			Gehan-Breslow-Wilcoxon test			Gehan-Breslow-Wilcoxon test		
Chi square	2.227		Chi square	5.825		Chi square	1.528	
df	1		df	1		df	1	
P value	0.1356		P value	0.0158		P value	0.2164	
P value summary	ns		P value summary	*		P value summary	ns	
Are the survival curves sig different?	No		Are the survival curves sig different?	Yes		Are the survival curves sig different?	No	
Median survival			Median survival			Median survival		
Control	50.50		Control	50.50		Treatment1	76.50	
Treatment1	76.50		Treatment2	182.0		Treatment2	182.0	
Ratio (and its reciprocal)	0.6601	1.515	Ratio (and its reciprocal)	0.2775	3.604	Ratio (and its reciprocal)	0.4203	2.379
95% CI of ratio	0.2804 to 1.554	0.6433 to 3.567	95% CI of ratio	0.1026 to 0.7503	1.333 to 9.745	95% CI of ratio	0.1528 to 1.157	0.8647 to 6.546
Hazard Ratio (Mantel-Haenszel)			Hazard Ratio (Mantel-Haenszel)			Hazard Ratio (Mantel-Haenszel)		
Ratio (and its reciprocal)	1.898	0.5270	Ratio (and its reciprocal)	3.642	0.2746	Ratio (and its reciprocal)	2.151	0.4649
95% CI of ratio	0.7443 to 4.838	0.2067 to 1.344	95% CI of ratio	1.306 to 10.16	0.09847 to 0.7658	95% CI of ratio	0.7843 to 5.899	0.1695 to 1.275
Hazard Ratio (logrank)			Hazard Ratio (logrank)			Hazard Ratio (logrank)		
Ratio (and its reciprocal)	1.720	0.5813	Ratio (and its reciprocal)	3.130	0.3195	Ratio (and its reciprocal)	2.084	0.4797
95% CI of ratio	0.7895 to 4.560	0.2193 to 1.267	95% CI of ratio	1.360 to 9.751	0.1026 to 0.7353	95% CI of ratio	0.8024 to 5.767	0.1734 to 1.246

- Step 3: Apply Bonferroni correction: $0.05/3=0.06$ or initial p-values*3

Comparing more than 2 groups



- At any point in time, hazard of dying in mice with lung infection is:
 - almost 2 times higher in the control than in the treatment 1 group ($p=0.54$)
 - 3.6 times higher in the control than in the treatment 1 group ($p=0.04$)

